

Master 2 M0 2018 – 2019

## Analyse des séries financières

Examen final, Février 2019

3h00, sans aucun document

1. Let  $(\varepsilon_t)_{t \in \mathbf{Z}}$  a sequence of centered independent and identically distributed random variables such as  $\mathbb{E}(\varepsilon_0^2) = 1$ . If it exists we consider a sequence  $(X_t)_{t \in \mathbf{Z}}$  such as:

$$X_t = \varepsilon_t \sigma_t \quad \text{with} \quad \sigma_t = a_0 + a_1 |X_{t-1}| + b_1 \sigma_{t-1} \quad \text{for any } t \in \mathbf{Z} \quad (1)$$

where  $(a_0, a_1, b_1) \in [0, \infty)^3$  are unknown parameters with  $b_1 > 0$ .

- (a) Prove that  $(\varepsilon_t)$  is a stationary time series.  
 (b) Denote  $\mathbf{R}^\infty$  the space of sequences of real numbers with finite number of non-zero real numbers. Consider  $F : \mathbf{R}^\infty \rightarrow \mathbf{R}$  be a measurable function on  $\mathbf{R}^\infty$ . Assume also that  $F$  is Lipchitzian on  $\mathbf{R}^\infty$ : there exists a sequence  $(\ell_i)_{i \in \mathbf{N}^*}$  of real non-negative numbers such that for any  $x = (x_i)_i, y = (y_i)_i \in \mathbf{R}^\infty$ ,

$$|F(x) - F(y)| \leq \sum_{i=1}^{\infty} \ell_i |x_i - y_i|, \quad \text{where} \quad \sum_{i=1}^{\infty} \ell_i < \infty.$$

For  $t \in \mathbf{Z}$ , prove that  $(Y_t^{(n)})_n$  where  $Y_t^{(n)} = F(\varepsilon_t, \varepsilon_{t-1}, \dots, \varepsilon_{t-n}, 0, 0, \dots)$  is a Cauchy sequence on  $\mathbb{L}^2$ . Deduce that  $(Y_t)_{t \in \mathbf{Z}}$ , where  $Y_t = F((\varepsilon_{t-k})_{k \in \mathbf{N}})$  for  $t \in \mathbf{Z}$ , is a stationary second order process.

- (c) Let  $\alpha(\cdot)$  be the function such as  $\alpha(x) = a_1|x| + b_1$ . Show that  $\mathbb{E}[|\log(\alpha(\varepsilon_0))|] < \infty$  using Jensen Inequality. Deduce also that for any  $t \in \mathbf{Z}$ :

$$\left( \prod_{i=1}^n \alpha(\varepsilon_{t-i}) \right)^{1/n} \xrightarrow[n \rightarrow +\infty]{a.s.} e^\gamma \quad \text{with} \quad \gamma = \mathbb{E}[\log(\alpha(\varepsilon_0))]. \quad (2)$$

- (d) Suppose  $(\sigma_t)_{t \in \mathbf{Z}}$  exists in (1). Then, establish that  $\sigma_t = a_0 + \alpha(\varepsilon_{t-1})\sigma_{t-1}$  for any  $t \in \mathbf{Z}$ . Deduce by recurrence that for any  $t \in \mathbf{Z}$  and  $k \in \mathbf{N}^*$ ,

$$\sigma_t = \beta_t(k) + \alpha(\varepsilon_{t-1}) \times \dots \times \alpha(\varepsilon_{t-k}) \sigma_{t-k} \quad \text{with} \quad \beta_t(k) = a_0 \left( 1 + \sum_{i=1}^{k-1} \prod_{j=1}^i \alpha(\varepsilon_{t-j}) \right).$$

- (e) Using (2) prove that if  $\gamma < 0$  then for any  $t \in \mathbf{Z}$ ,  $\beta_t(n) \xrightarrow[n \rightarrow +\infty]{a.s.} \beta_t$  with  $\beta_t$  a positive random variable and satisfies  $\beta_t = a_0 + \alpha(\varepsilon_{t-1})\beta_{t-1}$ . Consequently, write  $X_t = F(\varepsilon_t, \varepsilon_{t-1}, \dots)$  and conclude about the existence and stationarity of  $(X_t)_{t \in \mathbf{Z}}$ . What's happening if  $a_0 = 0$ ?  
 (f) Assume now  $\gamma < 0$  and  $b_1 < 1$ . Using an iterating decomposition, show for any  $t \in \mathbf{Z}$ :

$$\sigma_t = \frac{a_0}{1 - b_1} + a_1 \sum_{j=0}^{\infty} b_1^j |X_{t-j-1}|.$$

Deduce  $\mathbb{E}(X_t | X_{t-1}, X_{t-2}, \dots)$  and  $\text{var}(X_t | X_{t-1}, X_{t-2}, \dots)$ . Is  $(X_t)$  a conditionally heteroskedastic process?

- (g) Assume now that  $(X_1, \dots, X_N)$  is observed and let  $\theta = (a_0, a_1, b_1)$ . Prove that the quasi-maximum likelihood estimator  $\hat{\theta}$  of  $\theta$  is:

$$\hat{\theta} = \underset{\theta \in \Theta}{\operatorname{argmin}} \left\{ \log \left( \frac{a_0}{1-b_1} \right) + \frac{1}{2} \left( \frac{(1-b_1)X_1}{a_0} \right)^2 + \sum_{i=2}^N \log \left( \frac{a_0}{1-b_1} + a_1 \sum_{j=0}^{i-2} b_1^j |X_{i-j-1}| \right) + \frac{1}{2} \left( \frac{X_i}{\frac{a_0}{1-b_1} + a_1 \sum_{j=0}^{i-2} b_1^j |X_{i-j-1}|} \right)^2 \right\}$$

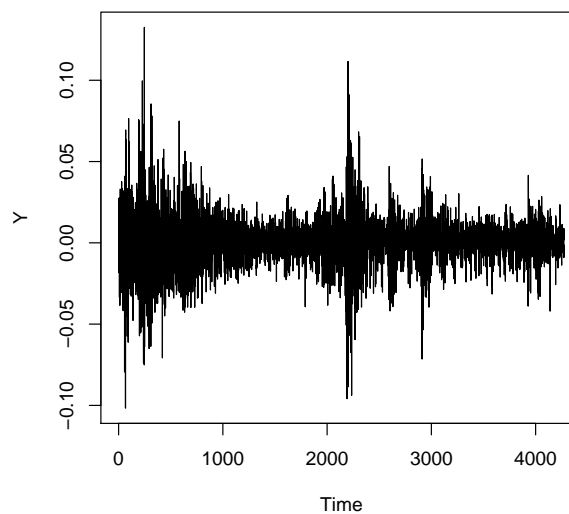
where  $\Theta$  is a set that should be specified.

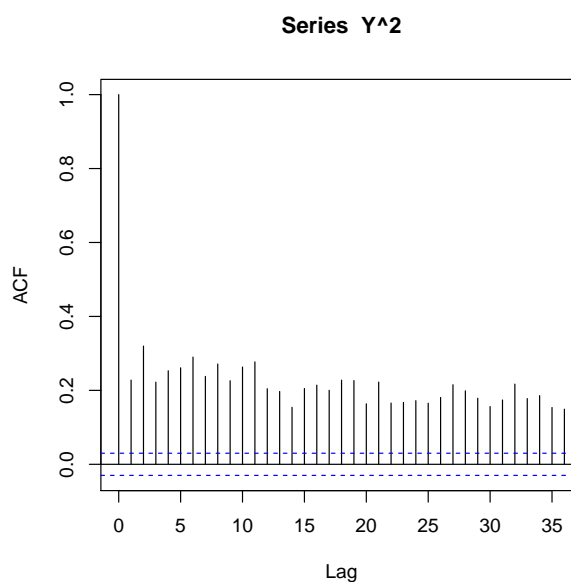
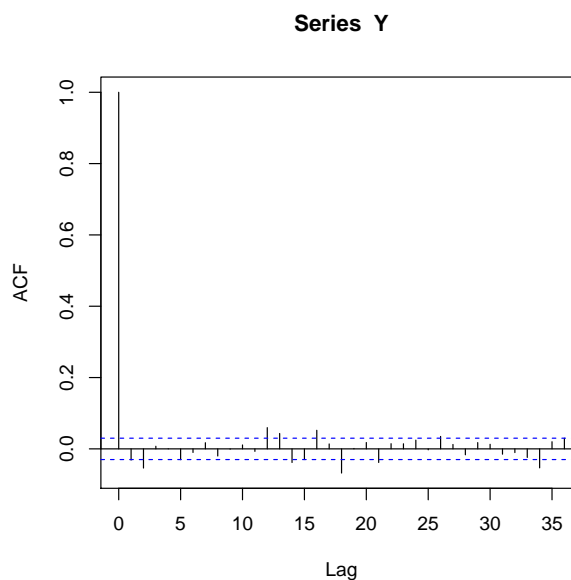
- (h) Is  $\hat{\theta}$  a consistent estimator? What is its convergence rate?  
 (i) Provide forecasting of  $X_{N+1}$  and  $X_{N+1}^2$ .

2. We study with R software the Nasdaq index from January 11st 2000 to January 11 2016.

- (a) First the following commands have been executed with figures exhibited below:

```
Nas=read.csv("C:/Users/Admin/Dropbox/Enseignement/M2 MMMEF/TP/Nasdaq.csv")
X=Nas$Closing
n=length(X)
Y=log(X[2:n]/X[1:(n-1)])
ts.plot(Y)
acf(Y)
acf(Y^2)
```





*Question II.1: Explain what is done. Which conclusions could you obtain from both the last commands? Are they compatible with a GARCH modelling?*

(b) New commands are then executed:

```
library(fGarch)
QMLE=garchFit(~ garch(1,1), data = Y, trace = FALSE)
QMLE
```

Here there are the numerical results:

Call:

```
garchFit(formula = ~garch(1, 1), data = Y, trace = FALSE)
```

Coefficient(s):

|  | mu         | omega      | alpha1     | beta1      |
|--|------------|------------|------------|------------|
|  | 6.4040e-04 | 2.1146e-06 | 8.6732e-02 | 9.0314e-01 |

|        | Estimate  | Std. Error | t value | Pr(> t ) |     |
|--------|-----------|------------|---------|----------|-----|
| mu     | 6.404e-04 | 1.617e-04  | 3.961   | 7.46e-05 | *** |
| omega  | 2.115e-06 | 4.052e-07  | 5.219   | 1.80e-07 | *** |
| alpha1 | 8.673e-02 | 8.639e-03  | 10.039  | < 2e-16  | *** |

```
beta1 9.031e-01 9.165e-03 98.544 < 2e-16 ***
```

|                   |                |                  | Statistic | p-Value      |
|-------------------|----------------|------------------|-----------|--------------|
| Jarque-Bera Test  | R              | Chi <sup>2</sup> | 168.5587  | 0            |
| Shapiro-Wilk Test | R              | W                | 0.9923722 | 2.300106e-14 |
| Ljung-Box Test    | R              | Q(10)            | 10.64455  | 0.3858725    |
| Ljung-Box Test    | R              | Q(15)            | 17.2698   | 0.3029932    |
| Ljung-Box Test    | R              | Q(20)            | 26.26701  | 0.1571734    |
| Ljung-Box Test    | R <sup>2</sup> | Q(10)            | 20.02718  | 0.02899662   |
| Ljung-Box Test    | R <sup>2</sup> | Q(15)            | 31.17497  | 0.008323147  |
| Ljung-Box Test    | R <sup>2</sup> | Q(20)            | 35.92566  | 0.01569342   |
| LM Arch Test      | R              | TR <sup>2</sup>  | 20.11987  | 0.06485208   |

Information Criterion Statistics:

| AIC       | BIC       | SIC       | HQIC      |
|-----------|-----------|-----------|-----------|
| -5.886147 | -5.880197 | -5.886149 | -5.884045 |

*Question II.2: Explain what is done and explain which conclusions you deduce.*

(c) Finally the following sequence of commands are executed:

```
M=matrix(0,3,4)
NAS10=garchFit(~ garch(1,0), data = Y, trace = FALSE)
M[1,1]=NAS10@fit$ics[2]
NAS11=garchFit(~ garch(1,1), data = Y, trace = FALSE)
M[1,2]=NAS11@fit$ics[2]
NAS12=garchFit(~ garch(1,2), data = Y, trace = FALSE)
M[1,3]=NAS12@fit$ics[2]
NAS13=garchFit(~ garch(1,3), data = Y, trace = FALSE)
M[1,4]=NAS13@fit$ics[2]
NAS20=garchFit(~ garch(2,0), data = Y, trace = FALSE)
M[2,1]=NAS20@fit$ics[2]
NAS21=garchFit(~ garch(2,1), data = Y, trace = FALSE)
M[2,2]=NAS21@fit$ics[2]
NAS22=garchFit(~ garch(2,2), data = Y, trace = FALSE)
.....
NAS33=garchFit(~ garch(3,3), data = Y, trace = FALSE)
M[3,4]=NAS33@fit$ics[2]
M
(Gop=which(M==min(M),2))
summary(NAS21)
```

The results are the following:

```
> M
      [,1]      [,2]      [,3]      [,4]
[1,] -5.501540 -5.880197 -5.878215 -5.876146
[2,] -5.665021 -5.881333 -5.879577 -5.877575
[3,] -5.717491 -5.879332 -5.877748 -5.875797
> (Gop=which(M==min(M),2))
      row col
[1,]  2   2
```

*Question II.3: What is done here and what are your conclusions?*